

MODELING COVARIATE ADJUSTED MORTALITY RELATIVE TO A
STANDARD POPULATION DOES BONE MARROW
TRANSPLANTATION PROVIDE A CLUE?

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chemotherapy. These studies, based on a Cox regression model, provide relative risk estimates of treatment modalities or prognostic indications. All estimates are relative to other patients in the disease.

With increasing follow-up of transplant patients it is natural to ask if bone marrow transplant in fact "cures" all patients or some subgroup of patients. Here, by "cured" meant patient's mortality rate is returned to the same mortality rate as would expect in a person of the same age and gender in the general population. While it is not reasonable to expect a return to the standard mortality rate of the general population immediately after transplant, it is possible that after some time the excess mortality directly related to the therapy may have abated out. Of interest is the estimation of the time of "cure" or the timing at a fixed time point to determine if the patient has been cured. It is also likely that this cure time may depend on some risk factors either known at the time of transplantation or by some point in time in the patients post transplant recovery process.

Twenty-five years ago the International Bone Marrow Transplant Registry (IBMTR) was founded with the goal of collecting data on consecutive allogeneic marrow transplants from multiple centers⁶. The IBMTR is a voluntary organization of 46 transplant teams worldwide that report all their consecutive cases to a central statistical center. Approximately 4% of the allogeneic transplants performed are reported to the Registry. Extensive data on patient risk factors is collected at the time of transplantation on most patients and patient follow-up information is obtained every six months.

In this note we present a model for the excess relative mortality due to transplantation in a group of 1,487 AML and 729 SAA patients from 14 countries. All patients included in the sample are alive and free of their primary disease at two years post transplant, so that all data observed in the sample are from causes not related to the sort term toxicity of the transplant itself. All patients were transplanted between 1981 and 1993. This is a subsample of a larger sample previously reported⁴ on which we are able to obtain current published life table information. Table 1 shows the distribution of the number of cases by the country where the patient was transplanted. Standard mortality tables were obtained for these countries by sex and for the US by sex and race (black versus non-black).

Of the 1,487 AML patients 16 died, while 34 of the 729 SAA patients died. For the AML patients the median follow-up was 6.2 years with a range of 2-16.7 years. For the aplastic anemia patients the median follow-up time was 6.7 years with a range of 2-16.8 years. The median age of the AML patients at the time of transplantation was 22.4 years (range 1.5-56.6 years) and was 18.8 years (range 0.2-69.4 years) for SAA patients.

There are a number of factors that are known to be predictive of survival following a transplant. One important factor is the development of graft-versus-host disease (GVHD). Two types of GVHD can occur, acute GVHD which occurs in the first 100 days post transplant and chronic GVHD which occurs after 100 days. We include as risk factors for survival a binary indicator of whether the patient had acute GVHD, an indicator of whether a patient had chronic GVHD prior to two years that was still active at two years, and indicator of whether a patient had chronic GVHD prior to two years that was resolved at two years. Age of the patient at the time of transplantation was also found to be

in transplant studies using the Cox model. While it is still of interest to see if young patients have a different “cur” rate than older patients. We divided the patients into three age groups: children (age ≤ 16 years), young patients (16-25 years) and older patients (> 25 years). A final covariate to be considered is the stage of the disease at the time of transplantation. For AML patients we classify patients as having early (transplanted in first complete remission), intermediate (transplanted in a second or later complete remission) or advanced (transplanted in relapse) disease. For SAA patients patients are classified as having early disease (time from diagnosis to transplant less than one year) or advanced disease (time from diagnosis to transplant more than one year). Table 2 summarizes the covariates for the disease.

To examine the effects of these covariates on survival the standard Cox regression model was fit to the data. For the model the hazard rate of an individual with covariate vector \mathbf{Z} is of the form

$$h(t|\mathbf{Z}) = h_0(t) \exp\{\boldsymbol{\gamma}^t \mathbf{Z}\}, \quad (1.1)$$

where $\boldsymbol{\gamma}$ is the vector of covariates and $h_0(t)$ is a baseline hazard rate. Here the risk coefficients, $\boldsymbol{\gamma}$, provide information on the relative effects of the covariates on survival among transplant patients and $h_0(t)$ is the default rate for, in our example, a child transplant patient with early disease of the type of GVHD. The results of fitting the standard Cox model are given in Table 3. These results show that for AML transplant patients, those with intermediate GVHD and intermediate disease tend to have lower survival, relative to other AML transplant patients. For SAA patients those with acute GVHD or active chronic GVHD and advanced disease, tend to have lower survival, relative to other SAA transplant patients.

In the next section we present a model for the survival of bone marrow transplant patients relative to the survival rates in the general population. The estimated relative mortality is allowed to be affected by a patient’s risk factors at the time of transplant. We develop a test of the hypothesis that the relative mortality is equal to one over a given time interval. This is a test that the mortality rate in the treated population over the interval is the same as that in the general population. In Section 3 we return to the example to determine at various times after transplant if a patient with a certain set of covariates has a mortality rate that has returned to normal.

2 A Model for

the fit of the model to the data

The data rate of the i th patient at t years post transplant is modelled as:

$$\lambda_i(t|\mathbf{Z}_i)$$

Applying Andersen's¹⁰ Corollary VII.2.6. if $Y_i(t)$ is replaced by $Y_i(t)\mu_i(t)$, it can be shown that a consistent estimator for the variance of $\hat{A}(s, t, \mathbf{Z}_0) = \hat{A}_0(s, t) \exp\{\hat{\beta}_t \mathbf{Z}_0$

From the tables compute the population mortality rate, $\lambda(a)$, at age a by assuming a constant mortality over the interval reported in the population life table. Under this assumption for an unbridged life table

$$\lambda(a) = -\ln[S(x+1)] - (-\ln[S(x)]), \quad \text{for } x \leq a < x+1,$$

if for a table with five year intervals compute

$$\lambda(a) = -\ln[S(x+5)] - (-\ln[S(x)])/5, \quad \text{for } x \leq a < x+1.$$

Once the population mortality rates are computed the value of $\mu_i(t)$ for a patient of age a_i at transplant is given by $\lambda(a_i+t)$, where $\lambda(\cdot)$ is from the proper age (race) and sex matched population. Using the population rates obtain the estimates of the relative mortality risk coefficients by maximizing (2.3). The estimates are given in Table 4.

An examination of Table 4 suggests that there is a significant effect of age on the relative mortality rate. Patients who are young are dying at a faster rate than older patients relative to the age matched mortality rates in the general population. Note that in the standard Cox model (Table 2), the comparisons are between transplant patients, there is no age effect for either disease. If there is no effect of age on transplant outcomes then the finding of an age effect in the relative mortality model is not surprising since younger patients have a lower population mortality rate. For both diseases the estimates of the effects of the other covariates are similar in the Cox model and the relative mortality model.

In Figures 1 and 2 plot a smoothed estimate of the relative mortality rate, $\hat{\lambda}_0(t) \exp(\hat{\beta}Z_0)$ for an AML and SAA patient in each of the two age groups. The plots are for patients who had not had graft-versus-host disease and were in the early disease state. The

For SAA patients the results presented in Table 6 show a different pattern. Here it appears that for patients over age 16 with no adverse risk factors the mortality rate is the same as in the general population after 4 years post transplant. For patients over age 25 with a single risk factor (active GVHD, prior history of acute GVHD or late disease) the mortality rate is the same as in the general population after 4 years, while if they have 2 or more risk factors the mortality rate is the same after 6 years. For young patients there is no difference between their mortality and the reference rates after 6 years if they have one of the risk factors presented.

4 Discussion

The techniques discussed here for estimation of the relative mortality rate are simple extensions of the Cox proportional hazards model. They are extended to include left truncated data by a simple redefinition of the risk set. The assumption of a proportional effect of the covariates on the relative mortality can be tested by using a time dependent covariate approach as in the usual proportional hazards regression model.

The test statistic (2.11) is a little poorer to detect a relative mortality rate increase on overall interval $[s, t]$. While it is mathematically possible that $\int_s^t \alpha_0(u)e^{\beta^t \mathbf{Z}_0} du = (t-s)$ and $\alpha_0(u)e^{\beta^t \mathbf{Z}_0} \neq 1$ for all $u \in [s, t]$, this would require that treated patients have a lower mortality rate than matched individuals in the general population. In most situations this is not biologically plausible.

As noted earlier the models above suggested by other authors and estimates of $A(s, t, \mathbf{Z}_0)$ are found in the papers. For this statistic the calculation of the variance of the estimator, requires some care since the estimator of $A(s, t, \mathbf{Z}_0)$ does not have independent increments.

In looking at the results in Tables 5 and 6 there is an obvious multiple testing problem in performing tests at different time points and at multiple covariate values. One could argue that some type of a corrected significance level should be used to make these comparisons of interest. We choose not to do so since our goal is to provide the investigator with only a crude notion of the patient's mortality rate as returned to normal and the p-values computed serve as measures of evidence against the hypothesis.

The ability to determine the relative mortality rate of a transplant recipient is a

insurance. This is currently a difficult and serious problem facing many transplant survivors.

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References

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Table 2. Frequencies of Covariates

COVA IATE	AML	SAA
Acut GVHD		
Y s	368 (24.7%)	145 (19.9%)
Non	1119 (75.3%)	584 (8 .1%)
C ronic Gv d		
Non	875 (58.8%)	465 (63.8%)
R solv d By 2 Y ars	236 (15.9%)	81 (11.1%)
Activ At 2 Y ars	376 (25.3%)	183 (25.1%)
Ag		
<16 Y ars	332 (22.4%)	284 (39. %)
16-25 Y ars	35 (23.5%)	251 (34.4%)
>25 Y ars	8 5 (54.1%)	194 (26.6%)
Dis as Stag		
Early	1132 (75.1%)	
Int rm diat	162 (1 .9%)	642 (88.1%)
Advanc d	193 (13. %)	87 (11.9%)

Table 3. Results Of Standard Cox Regression Analysis

Risk Factor	AML			SAA		
	$\hat{\beta}$	SE	p	$\hat{\beta}$	SE	p
Acute GVHD						
Yrs	.27	.176	.125	1.29	.349	.3
Chronic GVHD						
Resolved	.295	.224	.188	.592	.616	.337
Active	.398	.185	.32	1.468	.48	> .1
Age						
16-25	.141	.26	.588	-.84	.395	.831
>25	.438	.224	.5	.32	.424	.94
Dissemination						
Staged			< .1 ¹			
Intermediate	.67	.224	.7			
Advanced	.647	.2	.1	1.117	.38	.3

1. Test of goodness of fit: Wald test of effect of factor on survival.

Table . results Of relative Mortalit egression Anal sis

Risk Factor	AML			SAA		
	$\hat{\beta}$	SE	p	$\hat{\beta}$	SE	p
Acut GVHD						
Y s	.241	.175	.17	1.351	.396	< . 1
C ronic GVHD			. 678 ¹			. 3 ¹
R solv d	.3	.225	.182	.468	.626	.454
Activ	.414	.183	. 23	1.344	.4 7	. 1
Ag			< . 1 ¹			< . 1 ¹
16-25	- .716	.26	. 6	- .863	.395	. 29
>25	-1.339	.224	< . 1	-1.614	.426	< . 1
Dis as Stag			. 3 ¹			
Int rm diat	.666	.224	. 3			
Advanc d	.463	.2 1	. 21	1.168	.36	. 1

1. T o d gr of fr dom Wald t st of ff ct of factor on survival.

Table 5. p-Values Of The Test That The Mortality Rate For A Transplanted Patient Is The Same As In The General Population Over The Interval [s,12.6] For An AML Patient Without Acute GVHD

Age	Chronic GVHD	Dissemination	p-value n=8	p-value n=1
<16	Non	Early	.118	.2594
16-25	Non	Early	.37	.3917
>25	Non	Early	.1631	.636
<16	Activ	Early	.78	.2222
16-25	Activ	Early	.177	.316
>25	Activ	Early	.581	.457
<16	Non	Intermediate	.64	.27
16-25	Non	Intermediate	.125	.2655
>25	Non	Intermediate	.338	.3796
<16	Activ	Intermediate	.51	.1899
16-25	Activ	Intermediate	.81	.2259
>25	Activ	Intermediate	.116	.2943
<16	Non	Advanced	.75	.2188
16-25	Non	Advanced	.165	.2935
>25	Non	Advanced	.519	.4399
<16	Activ	Advanced	.57	.1973
16-25	Activ	Advanced	.98	.2428
>25	Activ	Advanced	.229	.336

Table 6. p-Values Of The Test That The Mortality Rate For A Transplanted Patient Is The Same As In The General Population Over The Interval [s,12.] For An Aplastic Anemia Patient

Age	Chronic GVHD	Disseminated State	Acute GVHD	p-value n s=2	p-value n s=4	p-value n s=6	p-value n s=8
<16	Non	Early	No	.11	.843	.3641	.4244
16-25	Non	Early	No	.1561	.7968	.9534	.927
>25	Non	Early	No	.9985	1.	1.	1.
<16	Active	Early	No	<.1	.51	.749	.1459
16-25	Active	Early	No	.1	.232	.181	.2623
>25	Active	Early	No	.48	.191	.5454	.5691
<16	Non	Late	No	<.1	.64	.859	.1597
16-25	Non	Late	No	.3	.359	.238	.393
>25	Non	Late	No	.133	.3195	.6865	.68
<16	Active	Late	No	<.1	.21	.44	.131
16-25	Active	Late	No	<.1	.37	.615	.1283
>25	Active	Late	No	<.1	.99	.117	.1888
<16	Non	Early	Yes	.39	.234	.982	.161
16-25	Non	Early	Yes	.12	.61	.254	.2736
>25	Non	Early	Yes	.481	.2453	.535	.562
<16	Active	Early	Yes	.23	.13	.611	.1151
16-25	Active	Early	Yes	.3	.174	.774	.136
>25	Active	Early	Yes	.49	.296	.118	.1836
<16	Non	Late	Yes	.23	.135	.632	.1178
16-25	Non	Late	Yes	.32	.191	.835	.1434
>25	Non	Late	Yes	.59	.354	.1359	.231
<16	Active	Late	Yes	.2	.112	.54	.155
16-25	Active	Late	Yes	.21	.123	.583	.1113
>25	Active	Late	Yes	.25	.147	.675	.1234

Relative Morta

